II. QUEUING THEORY

(a) General Concepts

- queuing theory useful for considering performance
    analysis of packet switching and circuit switching

General model of a queue:

- in practice, queue size is finite (i.e., number of packets
    that can be queued is limited → extra packets
    discarded → "blocking")

- if $\lambda > \mu \Rightarrow$ # queued packets will grow until queue
    saturated (remains full) or if queue size allowed to be
    ∞ (in theory), # queued packets will grow without
    bound

- $\rho = \lambda/\mu$ = utilization or traffic intensity

- as $\rho \rightarrow 1$, queue becomes unstable

- factors of interest: time delay, blocking performance, packet throughput (packets/time to get through)

- queue modelled by considering
     (1) packet arrival statistics
     (2) service time distribution (i.e., packet length distribution)
     (3) service discipline - FIFO, priority discipline
     (4) buffer size
     (5) input population (finite or ∞)

(b) Poisson Process

- arrival process (eg. packets generated at input to packet switch network or call initiated in circuit switch network) are often assumed to be Poisson

- continuous time:

- discrete time:

- divide time t into n intervals of length Δt (very small)

- let probability of arrival to queue in interval $\Delta t = p_+$ and assume all arrival events are independent (i.e., memoryless)

- assume Δt is small enough so that probability of ≥ 2 arrivals in Δt is negligible, i.e., $p_+ \ll 1$, then $p_+ \approx \lambda \Delta t$ (recall λ = arrival rate)

- rationale:

- average # of arrivals in an interval $t = \lambda t = np_+$

$$\Rightarrow p_+ = \lambda t/n$$

- probability of exactly k arrivals in $n = t/\Delta t$ intervals

$$P_k(n) =$$

(binomial distribution)

- hence,

- for fixed t, let $\Delta t \to 0 \Rightarrow n \to \infty$ since $t = n \cdot \Delta t$
  (i.e., making discrete case continuous)

$$\lim_{n \to \infty} P_k(n) = P_k(t)$$

(discrete)   (continuous)

- probability of k arrivals in a time t

$$P_k(t) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$$ Poisson Distribution

Notes:

(c) M/M/1 Queue

Interarrival Time

What is distribution?

- consider arbitrary point in time t_0 and define $t_0 = 0$

$$P(\text{arrival at time } t) = P(\text{no arrival in interval } (0,t))$$
$$\times P(\text{arrival in interval } (t, t + \Delta t))$$

- using independence

$$=$$

- consider graph of $f(t) = \lambda e^{-\lambda t}$

- since $f(\tau)\Delta t$ = probability = area under $f(t)$ then $f(t) = \lambda e^{-\lambda t}$
    is probability density function

- now $t_0 = 0$ can represent any arbitrary point in time, so
    since it can represent an arrival event point, the
    variable t represents an interarrival time

$\therefore$ interarrival time has exponential distribution with
    pdf $f(t) = \lambda e^{-\lambda t}$

Note:

Departures

- assume packets in queue and let p_- = probability of
    departure in interval Δt

- define $p_- = \mu \Delta t$ (recall μ = service or departure rate)

$\therefore$ P(departure after n intervals)

$$= \hspace{4cm} \text{(geometric distribution)}$$

$\therefore$ service/departure time pdf

$$f(t) = \mu e^{-\mu t} \hspace{1cm} \text{Exponential (same as arrivals)}$$

M/M/1 Queue:

M / *M* / 1
    $\rightarrow$ Markov Arrivals / Markov Departures / One Server

- Markov process $\rightarrow$ memoryless process

- *M/M*/1 $\Rightarrow$ Poisson arrivals, exponential service times, one
    server

(d) Discrete Model of M/M/1 Queue

- let k = # packets in queue including packet being served

- hence, k is a random variable and can be considered to be
    queue state

- now divide time into small intervals of Δt

State diagram:

- state transition for every interval

P_k = probability system in state k in an interval

$\Rightarrow$

Lemma 1

- by definition of pdf

Lemma 2

Theorem

- an interpretation

Proof of theorem by induction:

Base Case:

Induction Case:

- show if it is true for k - 1, it is true for k

Note:
- if you know one state probability and transition
    probabilities, you can determine probability of being
    in any state
- expect $P_k \to 0$ as $k \to \infty$ or queue will blow up

What is mean # of customers in queue?

$$\bar{k} = \frac{\rho}{1 - \rho}$$

What is variance of k? (variance is a measure of spread)

$$\sigma_k^2 = \frac{\rho}{(1-\rho)^2}$$

- conservation of customers for M/M/1 (∞ size)

What is time spent in queue?

Little's Theorem: $\bar{k} = \lambda \bar{T}$

where $\bar{T}$ = average time spent in system
(including service time)

intuition:

- if serviced in T and still $\bar{k}$ customers in queue, then for
  equilibrium $\bar{k}/\bar{T} = \lambda$

- makes intuitive sense but will not formally prove

- holds for M/M/1 and many other queues as well

(e) Queues with Finite Buffers

$M/M/1$ queue of size $N \rightarrow M/M/1/N$

- buffer overflow occurs when $k = N$ and packet arrives

- can use same state analysis as previous, except only $N+1$ states, instead of infinite number of states

- so

- mean of k

Conservation perspective:

but $\qquad \gamma = \mu(1\text{-}P_0)$

rate of servicing   fraction of time customer being served

$\therefore \; \mu(1\text{-}P_0) = \lambda(1\text{-}P_B)$

$$\rho = \frac{\lambda}{\mu} = \frac{1 - P_0}{1 - P_B}$$

$$\therefore \; P_B = \frac{P_0 - (1 - \rho)}{\rho}$$

- recall

$$\therefore \; P_B = P_N$$

(f) Extensions of M/M/1 Queue

(i) Multiple Sources

- combining two or more Poisson processes
  $\Rightarrow$ Poisson process

(ii) Multiple Servers

$M/M/m$ Queue:

- let *k* represent packets served and packets waiting and
    consider 2 cases:

(1) $k \leq m$ (i.e., all customers being served)

(2) $k \geq m$ (i.e., *m* customers being served, *k* - *m*
        waiting)

- for $k \leq m$

- for $k \geq m$

What is probability customer arrives at system and must wait to be served?

$$P_W =$$

$$P_W =$$

Erlang C Formula for M/M/m queue

(iii) Feedback

- simple communication system model

$\rightarrow$ in this case arrivals to Q2 are Poisson but loss may occur

- to minimize loss use feedback $\rightarrow$ feedback channel to shut off transmitter when receiver full

- arrivals toQ2 are now not Poisson, although $\mu_2 \geq \lambda_1$ or queue Q1 will blow up

(g) M/G/1 System

- often exponential service time is not an accurate model

    eg. in ATM, fixed size cell $\Rightarrow$ deterministic service time of cell size / link rate

- "G" represents general distribution for service time τ with known mean and variance

- let T = time in system, W = time waiting in queue

    so $T = W + \tau$

$$E\{W\} = \frac{\lambda E\{\tau^2\}}{2(1 - \rho)}$$

and average customers in queue given by $\bar{k} = \lambda \bar{T}$
    (Little's Theorem)

Special Cases:

Example Queuing Problem:

(h) Queuing Network Examples

- communication networks are, in fact, complex network of
   queues

Example 1:

Example 2:

Example 3:

Closed Queuing Networks

Aside: Norton equivalent of queuing network

N packets circulating around closed queuing network

- service rate dependent on number in queue

- derived by short circuitry A $\rightarrow$ B and allowing *n*
     customers to circulate

Example 4:

Sliding Window Flow Control with window size N

- assume all queues have same average service rate
    (i.e., same average packet sizes and link rates)

- assume ACKs are sent on high priority zero delay channel
    and are sent for every packet

- queue $M+1$ is an artificial queue used to represent
    generation of packets to send

- equivalent network:

What is $u(n)$?

∴ substituting (3) into (1) gives

$$P_n =$$

and then from (2)

$$P_0 =$$

- now throughput

$$\gamma =$$

- using Little's formula, average total delay

$\rightarrow$ could determine average delay from parameters N, M, μ, λ for sliding window flow control

- consider scenario where $\lambda \rightarrow \infty$ (i.e., packets served in zero time for queue $M+1$ implying data packets sent immediately following ACK)