

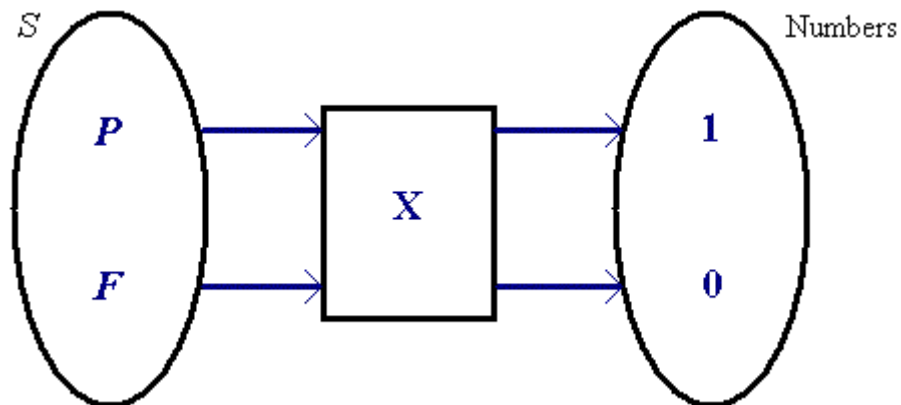
A **random quantity** [r.q.] maps an outcome to a number.

Example 6.01:

P = “A student passes ENGI 3423”

F = “That student fails ENGI 3423”

The sample space is $S = \{ P, F \}$



Define $X(P) = 1$, $X(F) = 0$, then
 X is a random quantity.

Definition: A **Bernoulli random quantity** has only two possible values:
 0 and 1.

Example 6.02

Let Y = the sum of the scores on two fair six-sided dice.

$$Y(i, j) = i + j$$

The possible values of Y are: $2, 3, 4, \dots, 12$

Example 6.03

Let N = the number of components tested when one fails.

The possible values of N are: $1, 2, 3, \dots$

A set D is **discrete** if

$n(D)$ is finite

OR

$n(D)$ is “countably infinite” (consecutive values can be found)

Examples:

6.03. Set \mathbb{N} = (the set of all natural numbers) is **discrete (countably infinite)**

6.04. $A = \{x : 1 \leq x \leq 2 \text{ and } x \text{ is real}\}$ is **not discrete (it is continuous)**

['1' is the smallest value, but what is the second-smallest value?]

A random quantity is **discrete** if its set of possible values is a discrete set.

Each value of a random quantity has some probability of occurring. The set of probabilities for all values of the random quantity defines a function $p(x)$, known as the

Probability Mass Function

(or probability function)

(*p.m.f.*):

$$p(x) = P[X = x]$$

Note: X is a random quantity, but x is a particular value of that random quantity.

All probability mass functions satisfy both of these conditions:

$$p(x) \geq 0 \quad \forall x$$

and

$$\sum_{\text{all } x} p(x) = 1$$

[Note that these two conditions together ensure that $p(x) \leq 1 \quad \forall x$.]

Example 6.05

$$f(x) = \begin{cases} cx^2 & x=1,2,3 \\ 0 & \text{otherwise} \end{cases} \leftarrow \text{[NOTE: may omit this branch]}$$

[$f(x) = 0$ is assumed for all x not mentioned in the definition of $f(x)$.]

$f(x)$ is a probability mass function. Find the value of the constant c .

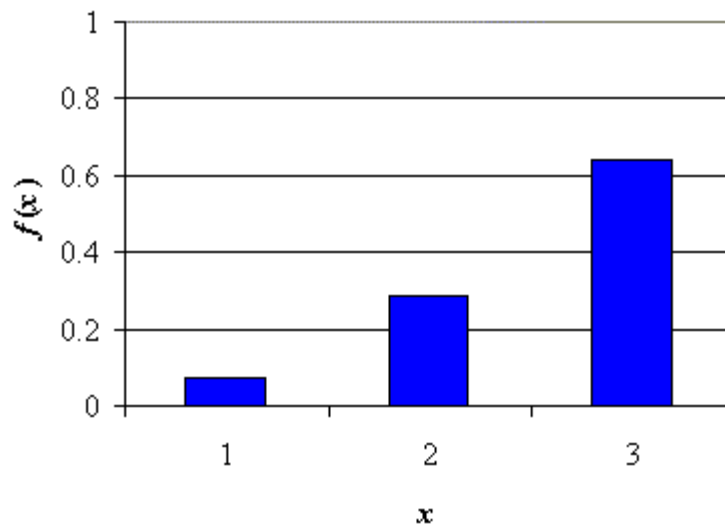
$$p(x) \geq 0 \quad \forall x \Rightarrow c \geq 0$$

$$\sum p(x) = 1 \Rightarrow c(1)^2 + c(2)^2 + c(3)^2 = 1$$

Therefore

$$c = \frac{1}{14}$$

Bar Chart:



Example 6.06

Find the p.m.f. for $X =$ (the number of heads when two fair coins are tossed).

Let $H_i =$ head on coin i and $T_i =$ tail on coin i .

The possible values of X are $X =$

0	(T_1T_2)
1	(H_1T_2 or T_1H_2)
or 2	(H_1H_2)

$$\begin{aligned}P[X = 0] &= P[T_1T_2] \\ &= P[T_1] P[T_2|T_1] \\ &= \frac{1}{2} \times \frac{1}{2} \quad (\text{independent events}) \\ &= \frac{1}{4}\end{aligned}$$

$$\begin{aligned}P[X = 1] &= P[H_1T_2] + P[T_1H_2] \quad (\text{mutually exclusive events}) \\ &= P[H_1] P[T_2] + P[T_1] P[H_2] \quad (\text{independent events}) \\ &= \frac{1}{2} \times \frac{1}{2} + \frac{1}{2} \times \frac{1}{2} \\ &= \frac{1}{4} + \frac{1}{4} \\ &= \frac{1}{2}\end{aligned}$$

$$\begin{aligned}P[X = 2] &= P[H_1H_2] \\ &= \frac{1}{2} \times \frac{1}{2} \\ &= \frac{1}{4}\end{aligned}$$

Therefore the p.m.f. is

$$f(x) = \begin{cases} \frac{1}{4} & (x = 0, 2) \\ \frac{1}{2} & (x = 1) \\ 0 & (\text{otherwise}) \end{cases}$$

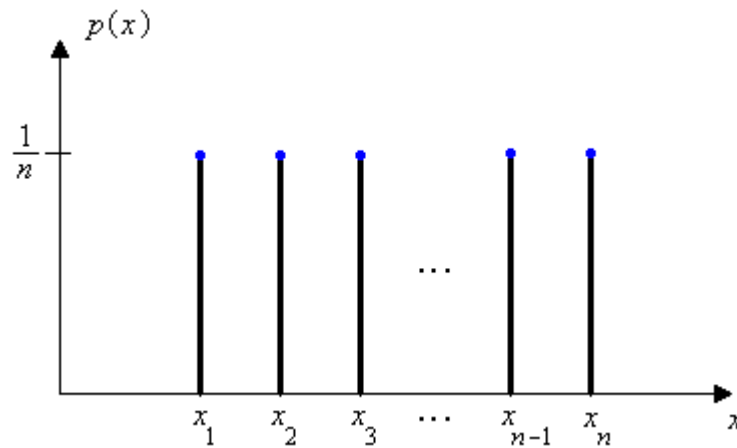
The Discrete Uniform Probability Distribution

A random quantity X , whose n possible values $\{x_1, x_2, x_3, \dots, x_n\}$ are all equally likely, possesses a discrete uniform probability distribution.

$$P[X = x_i] = \frac{1}{n} \quad (i = 1, 2, \dots, n)$$

An example is $X =$ (the score on a fair standard six-sided die), for which $n = 6$ and $x_i = i$.

Line graph:



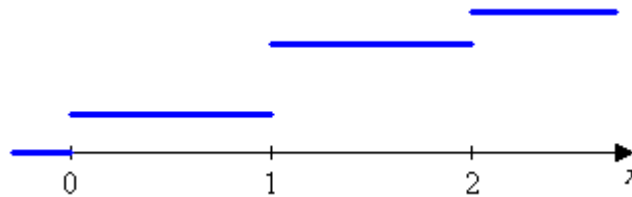
Cumulative Distribution Function (c.d.f.)

$$F(x) = P[X \leq x] = \sum_{y: y \leq x} p(y)$$

Example 6.07

Find the cumulative distribution function for
 $X =$ (the number of heads when two fair coins are tossed).

The possible values of X are 0, 1 and 2.



$$F(0) = P[X \leq 0] = p(0) = 1/4 .$$

$$\begin{aligned} F(1) &= P[X \leq 1] = P[X < 1] + P[X = 1] \\ &= F(0) + p(1) \\ &= 1/4 + 1/2 = 3/4 \end{aligned}$$

$$\begin{aligned} F(2) &= P[X \leq 2] \\ &= F(1) + p(2) \\ &= 3/4 + 1/4 = 1 \end{aligned}$$

$$\text{When } x < 0, \quad F(x) = P[X \leq x] \leq P[X < 0] = 0 \Rightarrow F(x) = 0$$

$$\text{When } x > 2, \quad F(x) = P[X \leq x] = F(2) + P[2 < X \leq x] = 1 + 0 = 1$$

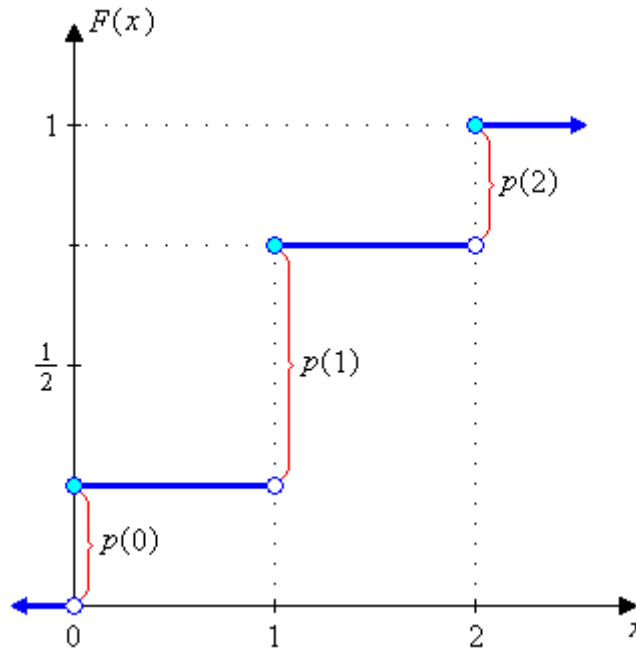
$$\text{When } 1 < x < 2, \quad P[X \leq x] = F(1) + P[1 < X \leq x] = 3/4 + 0 = 3/4$$

etc.

Thus

$$F(x) = \begin{cases} 0 & \text{if } x < 0 \\ 1/4 & \text{if } 0 \leq x < 1 \\ 3/4 & \text{if } 1 \leq x < 2 \\ 1 & \text{if } 2 \leq x \end{cases}$$

The graph of the c.d.f. is:



In general, the graph of a discrete c.d.f. :

- is always non-decreasing
- is level between consecutive possible values (staircase appearance)
- has a finite discontinuity at each possible value (step height = $p(x)$)
- rises in steps from $F(x) = 0$ to $F(x) = 1$.

Drawing convention:

- filled circle = point included in interval
- open circle = point excluded from interval

Example 6.08 (the inverse of the preceding problem):

Find the probability mass function $p(x)$ given the cumulative distribution function

$$F(x) = \begin{cases} 0 & \text{if } x < 0 \\ 1/4 & \text{if } 0 \leq x < 1 \\ 3/4 & \text{if } 1 \leq x < 2 \\ 1 & \text{if } 2 \leq x \end{cases}$$

Steps (= possible values) are at $x = 0, 1, 2$ only.

$$p(0) = F(0) = 1/4$$

$$p(1) = F(1) - F(0) = 3/4 - 1/4 = 1/2$$

$$p(2) = F(2) - F(1) = 1 - 3/4 = 1/4$$

and we recover the original p.m.f.



In general,

$$P[a < X \leq b] = F(b) - F(a)$$

last kept – last excluded

If a, b and **all** possible values are integers, then

$$P[a \leq X \leq b] = F(b) - F(a-1) \quad \text{and} \quad p(a) = P[X = a] = F(a) - F(a-1)$$

Example 6.09

Find and sketch the *c.d.f.* for $X =$ (the score upon rolling a fair standard die once).

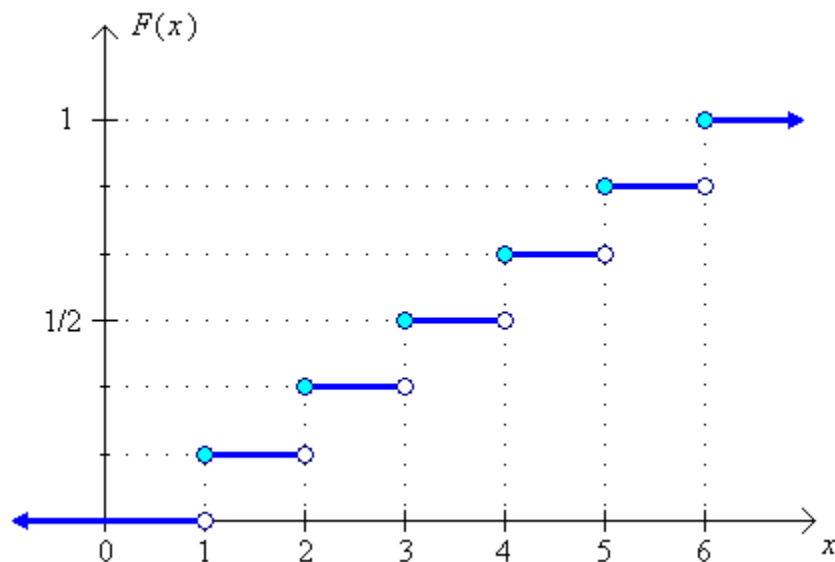
The *p.m.f.* is a uniform distribution

$$p(x) = \frac{1}{6} \quad (x = 1, 2, 3, 4, 5, 6)$$

Thus $F(x)$ increases from 0 to $1/6$ at $x = 1$ and increases by steps of $1/6$ at each subsequent integer value until $x = 6$. It follows easily that

$$F(x) = \begin{cases} 0 & (x < 1) \\ \text{INT}(x)/6 & (\text{otherwise}) \\ 1 & (x \geq 6) \end{cases}$$

The graph of $F(x)$ has the classic staircase appearance of the cumulative distribution function of a discrete random quantity.



Expected value of a random quantity

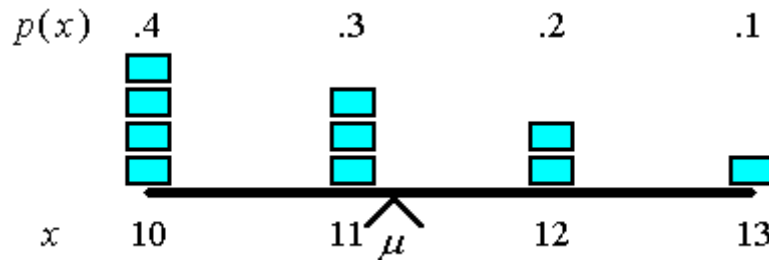
Example 6.10:

The random quantity X is known to have the p.m.f.

x	10	11	12	13
$p(x)$.4	.3	.2	.1

If we measure values for X many times, what value do we expect to see on average?

Treat the values of $p(x)$ as point masses of probability:



The **expected value** $E[X]$ (= **population mean** μ) is at the fulcrum (balance point) of the beam.

Taking moments about $x = 10$:

$$\sum p(x) (x-10) = .4 \times 0 + .3 \times 1 + .2 \times 2 + .1 \times 3 = 1.0$$

The fulcrum is at $x = 10 + 1$

Therefore $\mu = E[X] = \underline{11}$

In general, for any random quantity X with a discrete probability mass function $p(x)$ and a set of possible values D , the population mean μ of X (and the expected value of X) is

$$E[X] = \mu_x = \sum_{x \in D} x \cdot p(x)$$

Shortcut: If X is symmetric about $x = a$, then $E[X] = a$

Example 6.11:

Let X = the number of heads when a coin has been tossed twice. Find $E[X]$.

Solution:

List all the possible combinations.

→ the probability mass function of the distribution of X .

First toss	Second toss	X Number of heads
Tail	Tail	0
Tail	Head	1
Head	Tail	1
Head	Head	2

$$p(x) = \begin{cases} \frac{1}{4} & (x=0) \\ \frac{1}{2} & (x=1) \\ \frac{1}{4} & (x=2) \end{cases}$$

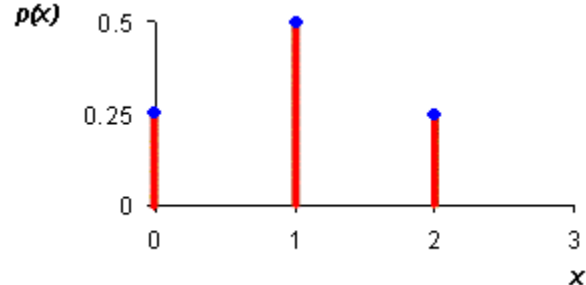
$$\begin{aligned} E[X] &= 0 \times \frac{1}{4} + 1 \times \frac{1}{2} + 2 \times \frac{1}{4} \\ &= 0 + \frac{1}{2} + \frac{1}{2} \end{aligned}$$

Therefore

$$\mu = \underline{1}$$

Alternative solution:

Graph of $p(x)$:



$p(x)$ is symmetric about $x = 1$.

Therefore, $E[X] = \underline{1}$

The expected value of a function

Definition:

If the random quantity X has set of possible values D and p.m.f. $p(x)$, then the expected value of any function $h(X)$, denoted by $E[h(X)]$, is computed by

$$E[h(X)] = \sum_{\text{all } x} h(x) \cdot p(x)$$

$E[h(X)]$ is computed in the same way that $E[X]$ itself is, except that $h(x)$ is substituted in place of x .

Special case:

$$h(x) = ax + b \Rightarrow \boxed{E[aX + b] = aE[X] + b}$$

Proof:

$$\begin{aligned} E[aX + b] &= \sum (ax + b) p(x) \\ &= \sum (ax) p(x) + \sum b p(x) \\ &= a \sum x p(x) + b \sum p(x) \\ &= a E[X] + b \end{aligned}$$

Example 6.12:

C = tomorrow's temperature high in °C

F = tomorrow's temperature high in °F

Given $E[C] = 10$, find $E[F]$.

$$F = \frac{9}{5}C + 32$$

$$\Rightarrow E[F] = \frac{9}{5}E[C] + 32$$

$$= \frac{9}{5} \times 10 + 32$$

$$\therefore E[F] = \underline{\underline{50}}$$

The variance of X

The quantity usually employed to measure the spread in the values of a random quantity

X is the **population variance** $V[X] = \sigma^2 = \frac{1}{N} \sum_x (x - \mu)^2$

Definition:

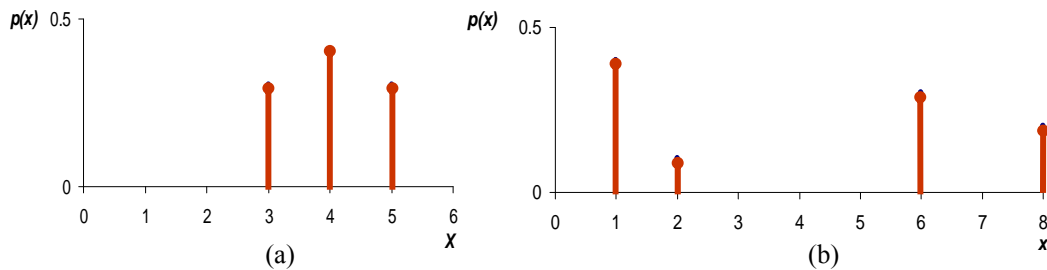
Let X have probability mass function $p(x)$ and expected value μ . Then

$$V[X] = \sum_x (x - \mu)^2 p(x) = E[(X - \mu)^2]$$

The **standard deviation** of X is $\sigma = \sqrt{V[X]}$

Example 6.13:

Two different probability distributions [below] share the same mean $\mu = 4$



If X has p.m.f. as shown in Figure (a)

x	3	4	5
$p(x)$.3	.4	.3

$\mu = 4$ (by symmetry)

$$V[X] = (3-4)^2 \times .3 + (4-4)^2 \times .4 + (5-4)^2 \times .3 = .3 + 0 + .3 = \underline{0.6}$$

and $\sigma = \sqrt{0.6} \approx 0.7746$

If X has p.m.f. as shown in Figure (b)

x	1	2	6	8
$p(x)$.4	.1	.3	.2

$$\mu = E[X] = 1 \times .4 + 2 \times .1 + 6 \times .3 + 8 \times .2 = .4 + .2 + 1.8 + 1.6 = \underline{4}$$

$$V[X] = (1-4)^2 \times .4 + (2-4)^2 \times .1 + (6-4)^2 \times .3 + (8-4)^2 \times .2 = 3.6 + 0.4 + 1.2 + 3.2 = \underline{8.4}$$

and $\sigma = \sqrt{8.4} \approx 2.898$ [Higher variance \leftrightarrow greater spread]

Example 6.14:

Let X = number of heads when a coin has been tossed twice. Find $V[X]$.

$$\begin{aligned} V[X] &= E[(X - \mu)^2] \\ &= \sum (x - \mu)^2 p(x) \\ &= (0-1)^2 \times \frac{1}{4} + (1-1)^2 \times \frac{1}{2} + (2-1)^2 \times \frac{1}{4} \\ &= \frac{1}{4} + 0 + \frac{1}{4} = \underline{0.5} \end{aligned}$$

A shortcut formula for variance

$$V[X] = E[X^2] - (E[X])^2$$

Proof: $\sigma^2 = E[X^2] - \mu^2$

$$\begin{aligned} V[X] &= E[(X - \mu)^2] = E[X^2 - 2\mu X + \mu^2] \\ &= E[X^2] - 2\mu E[X] + \mu^2 E[1] = E[X^2] - 2\mu\mu + \mu^2 \\ &= E[X^2] - \mu^2 \end{aligned}$$

Note: $E[f(X)] \neq f(E[X])$ unless $f(x)$ is linear and/or X is constant.

Example 6.14 (continued):

Let X = number of heads when a coin has been tossed twice. Find $V[X]$ using the shortcut formula.

$$\begin{aligned} E[X^2] &= \sum x^2 p(x) \\ &= 0^2 \times \frac{1}{4} + 1^2 \times \frac{1}{2} + 2^2 \times \frac{1}{4} \\ &= 0 + \frac{1}{2} + 1 = 1.5 \end{aligned}$$

$$\begin{aligned} V[X] &= E[X^2] - (E[X])^2 \\ &= 1.5 - 1^2 \end{aligned}$$

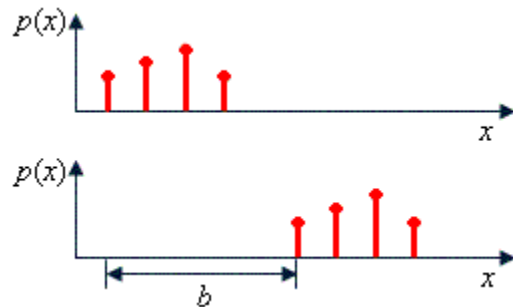
$$\therefore \sigma^2 = \underline{0.5}$$

The shortcut is more convenient when μ is not an integer.

Rules of variance

Example 6.15:

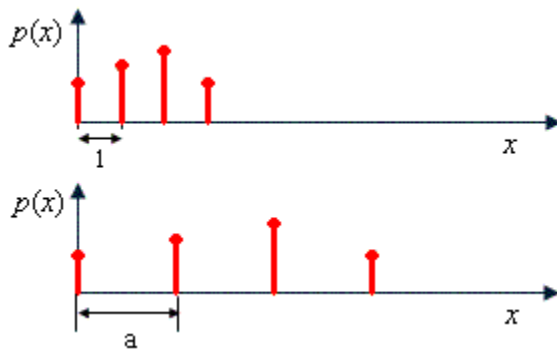
Do the distributions in the following two figures have the same variance or not?



YES

Example 6.16:

Do the distributions in the following two figures have the same variance or not?



NO

$$\boxed{V[aX + b] = a^2 V[X]}$$

Proof:

$$\begin{aligned} V[aX + b] &= E[(aX + b) - E[aX + b]]^2 \\ &= E[(aX + b) - (a\mu + b)]^2 \\ &= E[(aX - a\mu)]^2 \\ &= E[a^2 (X - \mu)^2] \\ &= a^2 E[(X - \mu)^2] \\ &= a^2 V[X] \end{aligned}$$

The addition of the constant b does not affect the variance, because the addition of b changes the location (and therefore mean value) but not the spread of values.