One simple test to determine whether or not a given random sample is consistent with some probability distribution is to compare the numbers of observations in each of $k$ intervals with the numbers that would be expected if the probability distribution is correct.

Let the observed values be $\left\{o_{1}, o_{2}, \ldots, o_{k}\right\}$.
The null hypothesis is $\mathscr{H}_{0}: p_{1}=p_{10}, p_{2}=p_{20}, \ldots, p_{k}=p_{k_{0}}$, (that is, the probability distribution is correct), where $p_{i}$ is the probability that the random quantity will fall in the $i^{\text {th }}$ interval.
The number of values expected in the $i^{\text {th }}$ interval when the null hypothesis is true is $e_{i}=n p_{i}$, where $n=\sum_{i=1}^{k} e_{i}=\sum_{i=1}^{k} O_{i}$ is the total number of observations.
The chi-square test statistic is

$$
\chi^{2}=\sum_{i=1}^{k} \frac{\left(O_{i}-e_{i}\right)^{2}}{e_{i}}
$$

Provided that all expected values are sufficiently large (about 5 or greater), this test statistic follows a chi-square distribution with $(k-1)$ degrees of freedom, to a good approximation.

Clearly, the closer all observed values are to the expected values, the lower the value of $\chi^{2}$ will be. If the test statistic exceeds $\chi_{\alpha, k-1}^{2}$, (the value of the chi-square distribution with $(k-1)$ degrees of freedom, above which $\alpha$ of the probability lies), then there is sufficient evidence to reject the null hypothesis in favour of the alternative hypothesis that the sample did not come from the hypothesized probability distribution.

The probability density function for $\chi_{5}^{2}$ is shown here:


## Example 14.01

A standard six-sided die is rolled 60 times, with results as shown. Can one conclude, at a $5 \%$ level of significance, that the die is loaded (that is, the six faces are not all equally likely)?

| Score: | 1 | 2 | 3 | 4 | 5 | 6 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |


| Score | $o_{i}$ | $e_{i}$ | $\left(o_{i}-e_{i}\right)$ | $\left(o_{i}-e_{i}\right)^{2}$ | $\left(o_{i}-e_{i}\right)^{2} / e_{i}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 8 |  |  |  |  |
| 2 | 9 |  |  |  |  |
| 3 | 7 |  |  |  |  |
| 4 | 10 |  |  |  |  |
| 5 | 8 |  |  |  |  |
| 6 | 18 |  |  |  |  |
|  |  |  |  |  |  |

The test statistic is
$\chi^{2}=\sum_{i=1}^{6} \frac{\left(o_{i}-e_{i}\right)^{2}}{e_{i}}=$
Compare the test statistic to the critical value

Note that the chi-square goodness-of-fit test is more complicated if the parameters of the probability distribution have to be estimated from the data. We shall not explore such a situation in this course.

Below is an extract from a chi-square table [Navidi Table A.6; Devore Table A.7].
The full table and a calculator for the chi-square distribution are available as Excel spreadsheet files at " www.engr.mun.ca/~ggeorge/4421/demos/".

|  | $\boldsymbol{\alpha}$ |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| $\boldsymbol{v}$ | $\mathbf{0 . 1}$ | $\mathbf{0 . 0 5}$ | $\mathbf{0 . 0 2 5}$ | $\mathbf{0 . 0 1}$ | $\mathbf{0 . 0 0 5}$ |
| $\mathbf{1}$ | 2.70554 | 3.84146 | 5.02389 | 6.63490 | 7.87944 |
| $\mathbf{2}$ | 4.60517 | 5.99146 | 7.37776 | 9.21034 | 10.59663 |
| $\mathbf{3}$ | 6.25139 | 7.81473 | 9.34840 | 11.34487 | 12.83816 |
| $\mathbf{4}$ | 7.77944 | 9.48773 | 11.14329 | 13.27670 | 14.86026 |
| $\mathbf{5}$ | 9.23636 | 11.07050 | 12.83250 | 15.08627 | 16.74960 |
| $\mathbf{6}$ | 10.64464 | 12.59159 | 14.44938 | 16.81189 | 18.54758 |
| $\mathbf{7}$ | 12.01704 | 14.06714 | 16.01276 | 18.47531 | 20.27774 |
| $\mathbf{8}$ | 13.36157 | 15.50731 | 17.53455 | 20.09024 | 21.95495 |
| $\mathbf{9}$ | 14.68366 | 16.91898 | 19.02277 | 21.66599 | 23.58935 |
| $\mathbf{1 0}$ | 15.98718 | 18.30704 | 20.48318 | 23.20925 | 25.18818 |
| $\mathbf{1 1}$ | 17.27501 | 19.67514 | 21.92005 | 24.72497 | 26.75685 |
| $\mathbf{1 2}$ | 18.54935 | 21.02607 | 23.33666 | 26.21697 | 28.29952 |
| $\mathbf{1 3}$ | 19.81193 | 22.36203 | 24.73560 | 27.68825 | 29.81947 |
| $\mathbf{1 4}$ | 21.06414 | 23.68479 | 26.11895 | 29.14124 | 31.31935 |
| $\mathbf{1 5}$ | 22.30713 | 24.99579 | 27.48839 | 30.57791 | 32.80132 |
| $\mathbf{1 6}$ | 23.54183 | 26.29623 | 28.84535 | 31.99993 | 34.26719 |
| $\mathbf{1 7}$ | 24.76904 | 27.58711 | 30.19101 | 33.40866 | 35.71847 |
| $\mathbf{1 8}$ | 25.98942 | 28.86930 | 31.52638 | 34.80531 | 37.15645 |
| $\mathbf{1 9}$ | 27.20357 | 30.14353 | 32.85233 | 36.19087 | 38.58226 |
| $\mathbf{2 0}$ | 28.41198 | 31.41043 | 34.16961 | 37.56623 | 39.99685 |

## The Chi-Square Test for Independence

## Example 14.02

The lengths of cables from a production line are classified as too short, OK and too long. The diameters of those cables are classified as too thin, OK and too thick.

Measurements of 300 cables produce the following table

|  |  | Diameter |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Too thin | OK | Too thick | Total |
| Length | Too short | 4 | 18 | 53 | 75 |
|  | OK | 32 | 77 | 26 | 135 |
|  | Too long | 64 | 20 | 6 | 90 |
|  | Total | 100 | 115 | 85 | 300 |

Can one conclude, at a $5 \%$ level of significance, that the diameters are not independent of the lengths of the cables?

Let the observed value in row $i$, column $j$ be labelled $o_{i j}$.
Let the sum of the values in row $i$ be labelled $o_{i \bullet}=\sum_{j=1}^{3} o_{i j}=o_{i 1}+o_{i 2}+o_{i 3}$
Let the sum of the values in column $j$ be labelled $o_{\cdot j}=\sum_{i=1}^{3} o_{i j}=o_{1 j}+o_{2 j}+o_{3 j}$
The grand total number of observations is
$n=o_{. .}=\sum_{j=1}^{3} o_{\bullet}{ }_{j}=\sum_{i=1}^{3} o_{\bullet}=\sum_{i=1}^{3} \sum_{j=1}^{3} o_{i j}=o_{11}+o_{12}+\ldots+o_{32}+o_{33}$
Recall that if and only if two events $A, B$ are independent, then $\mathrm{P}[A B]=\mathrm{P}[A] \times \mathrm{P}[B]$
The event "length too short" is independent of the event "diameter too thin" iff P ["length too short" $\bigcap$ "diameter too thin" $]=\mathrm{P}[$ "length too short" $] \times \mathrm{P}[$ "diameter too thin" $]$
This leads to
E [\# "length too short" $\cap$ "diameter too thin"] $=n \times \mathrm{P}$ ["length too short" $\cap$ "diameter too thin"]
$=n \times \frac{\#(\text { "length too short" })}{n} \times \frac{\#(\text { "diameter too thin" })}{n}$
$\Rightarrow \quad e_{11}=\frac{o_{1} \times o_{.1}}{n}=\frac{75 \times 100}{300}=25$
iff $O_{1 .}=$ "length too short" is independent of $O_{.1}=$ "diameter too thin"

## Example 14.02 (continued)

The other expected values are generated in the same way:

$$
e_{i j}=\frac{o_{i \cdot} \times o_{\cdot j}}{o_{. .}}
$$

We are testing, at $\alpha=5 \%$,
$\mathcal{H}_{0}$ : the diameter and length are independent
vs.
$\mathscr{H}_{\mathrm{A}}$ : the diameter and length are not independent

If $\mathscr{H}_{0}$ is true, then the expected values are:

|  |  | Diameter |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Length | Too thin | OK | Too thick | Total |  |
|  | Too short | 25 | 28.75 | 21.25 | 75 |
|  | OK | 45 | 51.75 | 38.25 | 135 |
|  | Too long | 30 | 34.50 | 25.5 | 90 |
|  | Total | 100 | 115 | 85 | 300 |

The test statistic is
$\chi^{2}=\sum_{i=1}^{3} \sum_{j=1}^{3} \frac{\left(o_{i j}-e_{i j}\right)^{2}}{e_{i j}}=\frac{(4-25)^{2}}{25}+\frac{(18-28.75)^{2}}{28.75}+\frac{(53-21.25)^{2}}{21.25}+\ldots+\frac{(6-25.50)^{2}}{25.50}$
Note that one observed value is less than 5, but that doesn't matter. What does matter is: all nine expected values are greater than 5 . Tedious calculation results in $\chi^{2}=148.6 \ldots$

The number of degrees of freedom is

Compare the test statistic to the critical value
[Space for Additional Notes]

